

Data based Quantitative Analysis and Calculation under Nonlinear Expectations

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AQFC 2017, April 24, Seoul

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- Covering probability uncertainty $\{P_\theta\}_{\theta \in \Theta}$
- Nonlinearity is fundamental and necessary
- Worst case philosophy:

$$\hat{\mathbb{E}}[X] := \max_{\theta \in \Theta} E_{P_\theta}[X] :$$

$\hat{\mathbb{E}}$ becomes an sublinear operation, called sublinear expectation

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- $\mathbb{E}[X_i] \downarrow 0$, if $X_i(\omega) \downarrow 0, \forall \omega$

Covering the risk of probability uncertainty by **sublinear expectation**

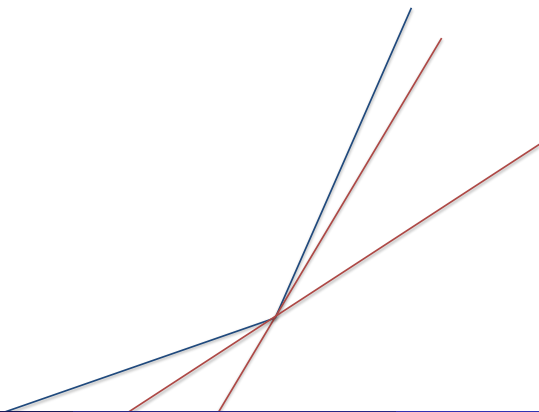
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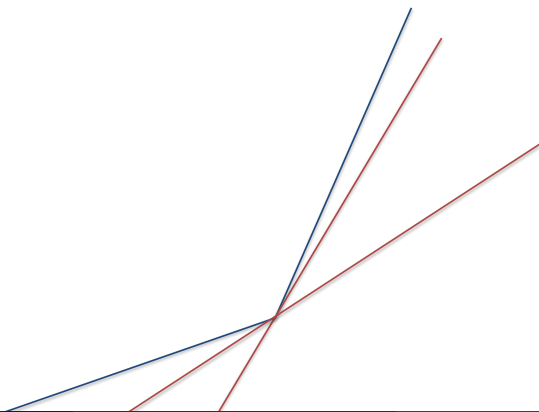
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Definition

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$$X \stackrel{d}{=} Y \iff \mathbb{E}[\varphi(X)] = \mathbb{E}[\varphi(Y)],$$

$$X \stackrel{d}{\geq} Y \iff \mathbb{E}[\varphi(X)] \geq \mathbb{E}[\varphi(Y)],$$

- Y is **independent** of X if

$$\mathbb{E}[\varphi(X, Y)] = \mathbb{E}[\mathbb{E}[\varphi(x, Y)]_{x=X}].$$

A toy but typical model: nonlinear Bernoulli model

- Randomly choose a ball from an urn (black box) containing Black and White balls ($W_1 + B_1 = 100$)
- But we know only $W_1 \in [40, 60]$

-

$$\tilde{\zeta}_1(\omega) = 1_{\{W_1 = \text{true}\}} - 1_{\{B_1 = \text{true}\}}$$

Repeat the above game: $i = 1, 2, 3, \dots$,

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- Each time i , W_i is changed within $W_i \in [40, 60]$
- We get a sequence of random variables $\{\xi_i(\omega)\}_{i=1}^{\infty}$;
- This is a typical case of our daily uncertainty: environment changes all the time
- The output $\{\xi_i\}_{i=1}^{\infty}$ is **i.i.d.**:
 - ξ_1, ξ_2, \dots are identically distributed (same distribution uncertainty)
 - ξ_{i+1} is independent of $\{\xi_i\}_{i=1}^n$.
- $S_n = \sum_{i=1}^n \xi_i$: a **nonlinear Bernoulli random walk**.

A general random generator of nonlinear i.i.d. sequence

$$\{\xi_i\}_{i=1}^{\infty}$$

- The 'black box' becomes a machine to output random vectors

$$\xi_i(\omega) \stackrel{d}{=} F_{\theta}, \quad i = 1, 2, 3, \dots$$

but we only know:

$$\mathcal{L}(\xi_i) \in \{F_{\theta}\}_{\theta \in \Theta}, \quad i = 1, 2, 3, \dots$$

A general random generator of nonlinear i.i.d. sequence

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- $\tilde{\zeta}_{i+1} \stackrel{d}{=} \tilde{\zeta}_i$, $\tilde{\zeta}_{i+1}$ is independent of $(\tilde{\zeta}_1, \tilde{\zeta}_2, \dots, \tilde{\zeta}_i)$
- Examp.: $\{F_{\theta}\}_{\theta \in \Theta} = M_{[\underline{\mu}, \bar{\mu}]}$: all prob. distributions concentrated on $[\underline{\mu}, \bar{\mu}]$.

Important case: i.i.d. with a maximal distribution

- $M_{[\underline{\mu}, \bar{\mu}]}$ -distributed i.i.d sequence $\{\tilde{\zeta}_i\}_{i=1}^{\infty}$: covers **all possible distributions of sequences** satisfying

$$\underline{\mu} \leq \tilde{\zeta}_i(\omega) \leq \bar{\mu}, \quad i = 1, 2, 3, \dots$$

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- Thus, principally, all random sequences can be treated as an nonlinear i.i.d. random sequence.
- Important task: how to narrow down $\bar{\mu} - \underline{\mu}$?
- Challenging objective: to establish a systematic framework $(\Omega, \mathcal{H}, \mathbb{E})$ compatible with (Ω, \mathcal{F}, P)
- Important: hidden behind are PDEs (linear/nonlinear heat equation)!

Two fundamentally important nonlinear distributions

$$\zeta_1 + \zeta_2 \stackrel{d}{=} 2\zeta_1 \iff \zeta_1 \stackrel{d}{=} M_{[\underline{\mu}, \bar{\mu}]}$$

$$\zeta_1 + \zeta_2 \stackrel{d}{=} \sqrt{2}\zeta_1 \iff \zeta_1 \stackrel{d}{=} N(0, [\underline{\sigma}^2, \bar{\sigma}^2])$$

Theorem (Peng2008-2010)

Let $\{Y_i\}_{i=1}^{\infty}$ be i.i.d. sequence. Assume $\mathbb{E}[|Y_1|^{1+\delta}] < \infty$.

Nonlinear Law of large number

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$$\lim_{n \rightarrow \infty} \mathbb{E}\left[\varphi\left(\frac{Y_1 + \cdots + Y_n}{n}\right)\right] = \mathbb{E}[(\varphi(Y))] = \max_{v \in [\underline{\mu}, \bar{\mu}]} \varphi(v).$$

where $\bar{\mu} = \mathbb{E}[Y_1]$, $\underline{\mu} = -\mathbb{E}[-Y_1]$.

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$Y \stackrel{d}{=} M_{[\bar{\mu}, \underline{\mu}]}$: Maximal distribution.

$u(x, t) := \mathbb{E}[\varphi(x + (1-t)Y)]$ solves the 1st order PDE:

$$\partial_t u(t, x) + g(\partial_x u) = 0, \quad u(1, x) = \varphi(x)$$

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Let $\{X_i\}_{i=1}^{\infty}$ be i.i.d. sequence. We assume furthermore that

$$\mathbb{E}[|X_1|^{2+\delta}] < \infty \quad \mathbb{E}[X_1] = \mathbb{E}[-X_1] = 0$$

Then, for each $\varphi \in C_b(\mathbb{R})$,

$$\lim_{n \rightarrow \infty} \mathbb{E}\left[\varphi\left(\frac{X_1 + \cdots + X_n}{\sqrt{n}}\right)\right] = \mathbb{E}[\varphi(X)].$$

where X is $N(0, [\underline{\sigma}^2, \bar{\sigma}^2])$ -distributed.

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$u(x, t) := \mathbb{E}[\varphi(x + \sqrt{1-t}X)]$ solves the 2nd order PDE

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A random variable X in $(\Omega, \mathcal{H}, \mathbb{E})$ is normal if the function

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where

- $G(a) = \frac{1}{2}[\bar{\sigma}^2 a^+ - \underline{\sigma}^2 a^-]$ $\bar{\sigma}^2 = \mathbb{E}[X^2]$, $\underline{\sigma}^2 = -\mathbb{E}[-X^2]$

A General misunderstanding caused from

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Classical 'Monté-Carlo' approach for estimating $\hat{\mathbb{E}}[\varphi(X)]$ through data

- Key point: How to obtain $\hat{\mathbb{E}}[\varphi(X)]$ through its sample $\{x_i\}_{i=1}^N$?
- In many practice cases: we care about $\hat{\mathbb{E}}[\varphi(X)]$ with a specific function $\varphi(x)$:
a consumption utility function, a contract, a cost function
- In a classical probability space (Ω, \mathcal{F}, P) , we can apply LLN to calculate

$$E[\varphi(X)] \sim \mathbb{M}[\varphi(X)] := \frac{1}{N} \sum_{i=1}^N \varphi(x_i)$$

where $\{x_i\}_{i=1}^N$ is an i.i.d. sample of X .

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- But: Is $\{x_i\}_{i=1}^N$ a classical **i.i.d.**?

- Convergence to a self-normalized G-Brownian motion, Zhengyan Lin and Zhixin Zhang, in PUQR, 2017.
- Nutz & van Handel (2013) Constructing sublinear expectations on path space. SPA, 123(8)
- P. 2006: G-expectation, G-Brownian motion and related stochastic calculus of Ito's type. The Abel Symposium 2005, Springer.
- P. (2008) A new central limit theorem under sublinear expectations
- P. (2008) Survey Sci. China Ser. A 52(7), 1391–1411 (2009)

- P. (2010) Nonlinear Expectations and Stochastic Calculus under Uncertainty (2010a). Preprint: arXiv:1002.4546 [math.PR]
- P. (2010) Tightness, weak compactness of nonlinear expectation (arXiv),
- Yan & Soner (2012) Weak approximation of G-expectations SPA,
- Zhang (2015) Donsker's invariance principle under the sublinear expectation with an application to Chung's law of the iterated logarithm. Commun. Math. Stat .
- Zhang,: (2016) Rosenthal's inequalities for independent and negatively dependent random variables under sub-linear expectations with applications. Sci. China Math.
- Fang X., Peng, S. & Shao Q. working paper

φ -max-mean algorithm: the data-based distribution of X

- Let $(\Omega, \mathcal{H}, \hat{\mathbb{E}})$ be a sublinear expectation space and

$$\{x_i\}_{i=1}^{n \times m} : \text{ i.i.d. sample of a random vector } X$$

- The max-mean algorithm to estimate $\hat{\mathbb{E}}[\varphi(X)]$:

$$\hat{\mathbb{M}}[\varphi] = \max\{Y_n^k : k = 1, \dots, m\},$$

where

$$Y_n^k = \frac{1}{n} \sum_{i=1}^n \varphi(x_{n(k-1)+i}).$$

- By nonlinear LLN, when $n \rightarrow \infty$, $\{Y_n^k\}_{k=1}^m \xrightarrow{d}$ an i.i.d. $\{Y^k\}_{k=1}^m$,

$$Y^k \stackrel{d}{=} M([\underline{\mu}_{\varphi(X)}, \bar{\mu}_{\varphi(X)}]).$$

- But $\max\{Y^k : k = 1, \dots, m\}$ provides us the asymptotically optimal unbiased estimate of $\bar{\mu}_{\varphi(X)} = \hat{\mathbb{E}}[\varphi(X)]$.

Optimality of the estimate

The optimality of the above estimate is based on the following quite simple, but very fundamental result:

Theorem (Jin-Peng2016)

Let Y^1, \dots, Y^m be i.i.d. and maximally distributed:

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$$\underline{\mu} \leq \min\{Y^1(\omega), \dots, Y^n(\omega)\} \leq \max\{Y^1(\omega), \dots, Y^n(\omega)\} \leq \bar{\mu}.$$

Moreover

$$\hat{\bar{\mu}}_n = \max\{Y^1, \dots, Y^n\},$$

is *the maximum unbiased estimate of $\bar{\mu}$.*

- Many typical nonlinear distributions

$$M_{[\underline{\mu}, \bar{\mu}]}, \quad N(\mu, [\underline{\sigma}^2, \bar{\sigma}^2]), \quad P_{[\underline{\lambda}, \bar{\lambda}]} \text{ (Nonlinear Poisson)}$$

- asymptotically unbiased estimates: ,

$$\hat{\underline{\sigma}}^2 := \min_{1 \leq k \leq m} \sigma_k^2, \quad \hat{\bar{\sigma}}^2 := \max_{1 \leq k \leq m} \sigma_k^2$$

$$\text{where } \sigma_k^2 := \frac{1}{n} \sum_{j=1}^n (x_{n(k-1)+j} - \mu)^2.$$

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Theorem.

If $(B_t)_{t \geq 0}$ is a G -Brownian motion and $\mathbb{E}[B_t] = \mathbb{E}[-B_t] \equiv 0$ then:

$$B_{t+s} - B_s \stackrel{d}{=} N(0, [\underline{\sigma}^2 t, \bar{\sigma}^2 t]), \forall s, t \geq 0$$



Real case study: from VaR to GVaR

Problem challenged by CFFEX

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$$\begin{aligned}\text{VaR}_\alpha^F(X) &= -\inf\{x \mid P(X \leq x) > \alpha\} \\ &= -\inf\{x \mid F(x) > \alpha\},\end{aligned}$$

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$$\begin{aligned}\text{VaR}_\alpha^F(X) &= -\inf\{x \mid P(X \leq x) > \alpha\} \\ &= -\inf\{x \mid F(x) > \alpha\},\end{aligned}$$

Can we use G -normal distribution in the place of a linear distribution F ?

- Nonlinear normally distributed VaR — **G-VaR**:

$$X_{t+1} \stackrel{d}{=} N(0, [\underline{\sigma}^2, \bar{\sigma}^2])$$

Real case study: from VaR to GVaR

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Empirical test of robust VaR

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F_G has the explicit expression:

$$F_G(x) = \int_{-\infty}^x \frac{\sqrt{2}}{\sqrt{\pi(\bar{\sigma} + \underline{\sigma})^2}} \left[\exp\left(\frac{-y^2}{2\bar{\sigma}^2}\right) \mathbf{1}_{y \leq 0} + \exp\left(\frac{-y^2}{2\underline{\sigma}^2}\right) \mathbf{1}_{y > 0} \right] dy. \quad (3)$$

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- For each \bar{t} , use the passed 1 year data $\{X_{\bar{t}-s}\}_{0 \leq s \leq l-1}$ to estimate two parameters $\underline{\sigma}_{\bar{t}}^2$ and $\bar{\sigma}_{\bar{t}}^2$ at the day \bar{t} :

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- Fix a window width $w = 100$ use the moving window

$$\sigma_{\bar{t},w}^2 := \sigma^2(X_{\bar{t}-w+1}, \dots, X_{\bar{t}}).$$

- Then get the upper and low data variances:

$$\bar{\sigma}_t^2 = \max\{\sigma_{t,20}^2, \sigma_{t-s,w}^2; s \in [0, \dots, l-w]\},$$

$$\underline{\sigma}_t^2 = \min\{\sigma_{t,20}^2, \sigma_{t-s,w}^2; s \in [0, \dots, l-w]\}.$$

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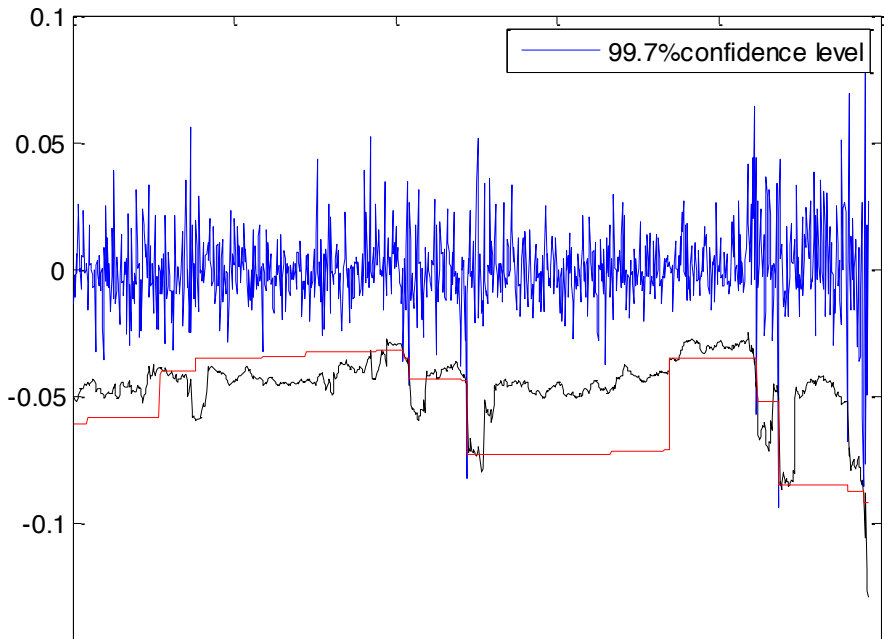
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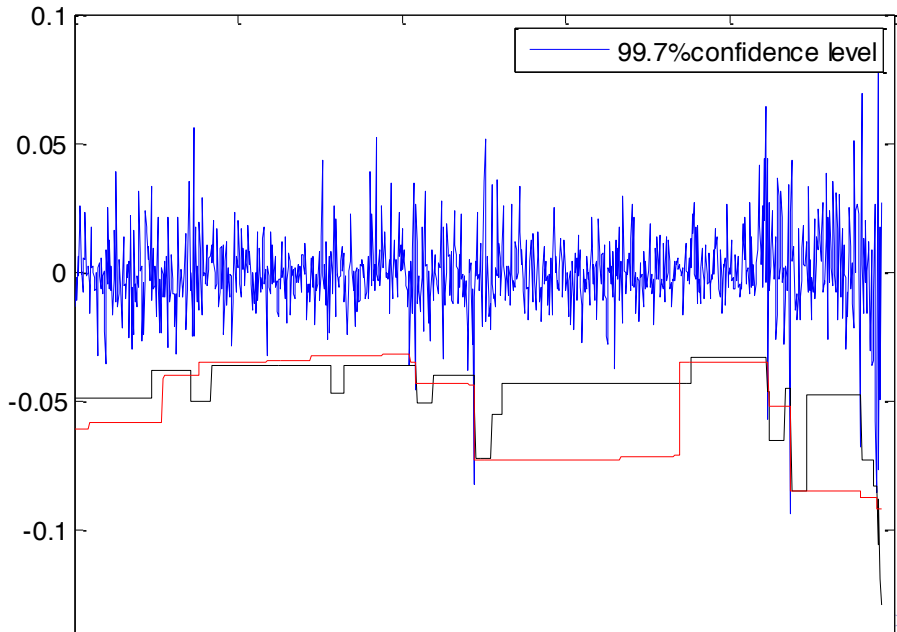
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- for uncertainty quantifications under probability-distribution uncertainties.

Probability v.s. Nonlinear Expectation

Probability Space	Nonlinear Expectation Space
(Ω, \mathcal{F}, P)	$(\Omega, \mathcal{H}, \mathbb{E})$: (sublinear is basic)
Distributions: $X \stackrel{d}{=} Y$	
Independence: Y indep. of X	
LLN and CLT	
Normal distributions	
Brownian motion $B_t(\omega) = \omega_t$	
Quadratic variable. $\langle B \rangle_t = t$	
Lévy process	

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Lévy process	nonlinear Lévy process

Probability v.s. Nonlinear Expectation

Probability Space	Nonlinear Expectation Space
Itô's calculus for BM	Itô's calculus for nonlinear BM
SDE $dx_t = b(x_t)dt + \sigma(x_t)dB_t$	
Diffusion: $\partial_t u - \mathcal{L}u = 0$	
Markovian pro. and semi-group	
Martingales	
$E[X \mathcal{F}_t] = E[X] + \int_0^T z_s dB_s$	

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Markovian pro. and semi-grou	Nonlinear Markovian
Martingales	Nonlinear Martingales
$E[X \mathcal{F}_t] = E[X] + \int_0^t z_s dB_s$	$\mathbb{E}[X \mathcal{F}_t] = \mathbb{E}[X] + \int_0^t z_s dB_s + K_t$

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Martingales	Nonlinear Martingales
$E[X \mathcal{F}_t] = E[X] + \int_0^t z_s dB_s$	$\mathbb{E}[X \mathcal{F}_t] = \mathbb{E}[X] + \int_0^t z_s dB_s + K_t$ $K_t = \int_0^t \eta_s d\langle B \rangle_s - \int_0^t 2G(\eta_s) ds$

New journal : Probability, Uncertainty and Quantitative Risk

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Thank you 谢谢